

# False confidence, non-additive beliefs, and valid statistical inference<sup>12</sup>

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ACMS Department Colloquium

University of Notre Dame

February 8th, 2019

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<sup>1</sup>Paper available at <https://www.researchers.one/article/2019-02-1>

<sup>2</sup>Thanks to friends and collaborators M. Balch, H. Crane, C. Liu...

- Statistics has developed a lot in the last 50+ years.
- But important fundamental questions about probability and statistical inference remain unanswered.
- Does it matter? YES!
- Currently we have two dominant schools of thought:
  - *frequentist*
  - *Bayesian*
- Very different views leading to different answers.
- Lots of debate over the years about which one is “right.”
- I think it's better to ask:
  - what does science need from statistics?
  - do existing approaches meet this need?
  - if not, then how to fill the void?

- Statistical problem:
  - Observable data is  $Y$ ;
  - Model is  $\mathcal{M} = \{P_{Y|\theta} : \theta \in \Theta\}$  — taken as given.
- Scientific questions correspond to hypotheses about  $\theta$ .
- Goal is to quantify uncertainty about  $\theta$ , given  $Y = y$ .
- Define an *inferential model*

$$(y, \mathcal{M}, \dots) \mapsto b_y : 2^\Theta \rightarrow [0, 1],$$

where  $b_y(A)$  represents the data analyst's degrees of belief about a hypothesis  $A \subseteq \Theta$  based on data  $y$ , model  $\mathcal{M}$ , etc.

- The inferential model could be lots of things:
  - a Bayesian posterior distribution;
  - a fiducial or confidence distribution;
  - ...

- Pointless if the inferential model isn't "reliable."
- Brad Efron said (roughly) that the construction of reliable, prior-free, inferential models is *the most important unresolved problem in statistical inference*.
- Key insight: *go non-additive!*
  - Familiar things are *additive*, i.e.,  $b_y$  is a probability.
  - But additivity isn't necessary, might even be a constraint.
  - "There's more to uncertainty than probabilities"<sup>3</sup>
- Take-away messages:
  - additive beliefs are afflicted with false confidence
  - good non-additive beliefs can avoid false confidence
  - there's a way to construct good non-additive beliefs

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<sup>3</sup><http://www.isipta2019.ugent.be>

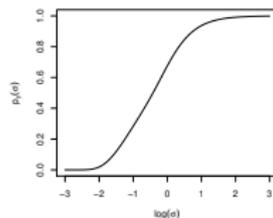
- The price of additivity:
  - satellite collision example
  - false confidence theorem
- Going non-additive:
  - avoiding false confidence: the validity property
  - non-additive beliefs and random sets
  - construction of valid inferential models
- A few technical remarks:
  - efficiency and dimension reduction
  - a complete-class theorem
- Conclusion

# Satellite collision problem

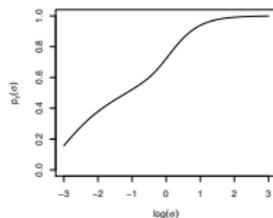
- A satellite orbiting Earth could collide with another object.
- Potential mess, so navigators try to avoid collision.
- Data on position, velocity, etc is used, along with physical and statistical models, to compute a *collision probability*.
- If collision probability is low, then satellite is judged safe; otherwise, some action is taken.
- An unusual phenomenon has been observed: noisier data necessarily makes collision probability small...

## ■ Illustration:

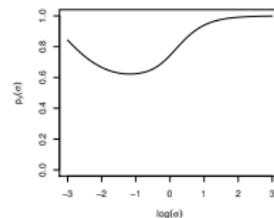
- $\|y\|$  denotes measured distance between satellite and object.
  - True distance  $\leq 1$  implies collision.
  - Measurement error variance is  $\sigma^2$ .
  - $p_y(\sigma)$  = probability of non-collision, given  $y$ .
- When  $\sigma$  is large,  $p_y(\sigma)$  is large, no matter what  $y$  is!
- *Potentially misleading conclusions!*



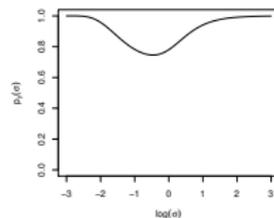
(a)  $\|y\|^2 = 0.5$



(b)  $\|y\|^2 = 0.9$



(c)  $\|y\|^2 = 1.1$



(d)  $\|y\|^2 = 1.5$

- What's going on here?
- Apparently, data are not sufficiently informative with respect to questions about collision/non-collision.
- *Additivity* forces the probability to go somewhere, happens that it goes to non-collision no matter what data says.
- Even if the satellite is on a direct collision course, probability tells navigators that it's safe.
- *False confidence*:<sup>4</sup> a hypothesis tending to be assigned large probability even though data does not support it.
- Is this a general phenomenon, or just something weird about this particular problem?

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<sup>4</sup>Balch, M., and Ferson, arXiv:1706.08565

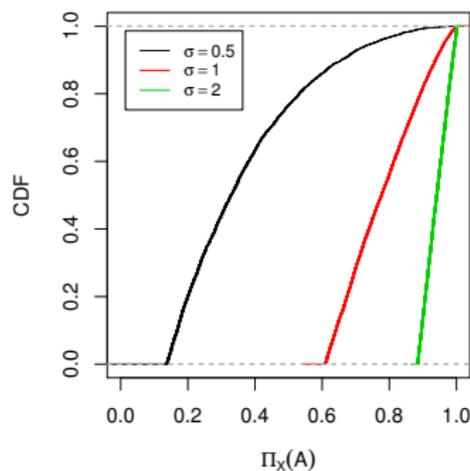
## False confidence theorem.

Let  $\Pi_Y$  be any additive belief on  $\Theta$ , depending on data  $Y$ . Then, for any  $\alpha$  and any  $t$ , there exists  $A \subset \Theta$  such that

$$A \not\subseteq \theta \quad \text{and} \quad P_{Y|\theta}\{\Pi_Y(A) > t\} \geq \alpha.$$

- In words, there always exists a false hypothesis that tends to be assigned high posterior probability.
- If judgments about the plausibility of a hypothesis are made based on the magnitude of its probability, then the theorem says there's risk of systematic error.
- Reveals a practical price attached to additivity.
- Doesn't say which hypotheses are afflicted, or to what extent, only that they exist.

- Simple version of the satellite example.
  - Let  $A$  denote the non-collision hypothesis.
  - Then  $\Pi_Y(A)$  as a random variable, with a CDF
  - Plot CDF when data are generated under a collision course.
- *False confidence*:  $\Pi_Y(A)$  is almost always large!



- Theorem says *every additive belief function* is afflicted with false confidence.
- Reid & Cox write:  
*it is unacceptable if a procedure . . . of representing uncertain knowledge would, if used repeatedly, give systematically misleading conclusions*
- To avoid false confidence and “systematically misleading conclusions,”  $b_y$  must be non-additive.
- e.g.,  $b_y(A) + b_y(A^c) < 1$ .
- But not every non-additive belief avoids false confidence.
- So we need some additional restrictions...

- First: *Why should we care about false confidence?*
  - Lots of Bayesian success stories, am I'm over-reacting?
  - e.g., have there been satellite collisions?
- The theorem doesn't say that something *will* go wrong, only that there's a risk.
- Statisticians directly involved in the data analysis might be able to manage their risk reasonably well.
- But this level of involvement is rare — we are largely focused on developing methods/software to be used by others.
- To manage risk under this increased exposure, I must either
  - inform users which hypotheses are too dangerous,
  - or avoid false confidence altogether.

## Definition.

An inferential model  $(y, \mathcal{M}, \dots) \mapsto b_y$  is *valid* if

$$\sup_{\theta \notin A} P_{Y|\theta} \{b_Y(A) > 1 - \alpha\} \leq \alpha, \quad \forall A \subseteq \Theta, \quad \forall \alpha \in (0, 1).$$

- Validity property implies that assigning high belief to a false hypothesis is a rare event — *no false confidence*.
- $\alpha$  on inside and outside calibrates the belief function values, i.e., so that I know what “large” and “small” means.
- Implies that procedures derived from a valid  $b_y$  have frequentist error rate control; details later.

- Simplest way to construct non-additive beliefs on a space  $\mathbb{X}$  is via a *random set*  $\mathcal{X} \sim P_{\mathcal{X}}$  that takes values in  $2^{\mathbb{X}}$ .
- For a fixed set  $A \subseteq \mathbb{X}$ , a realization of  $\mathcal{X}$  could be
  - fully contained in  $A$ ,
  - fully contained in  $A^c$ ,
  - or have non-empty intersection with both.
- Then the containment functional

$$b(A) = P_{\mathcal{X}}(\mathcal{X} \subseteq A), \quad A \subseteq \mathbb{X},$$

is a continuous, completely monotone Choquet capacity and, in particular,  $b(A) + b(A^c) \leq 1$ .

- Express the statistical model,  $P_{Y|\theta}$ , as

$$Y = a(\theta, U), \quad U \sim P_U, \quad P_U \text{ known.}$$

- Intuition: *If  $U$  were observable*, then just solve for  $\theta$  in terms of  $Y$  and  $U$  — done!
- Unfortunately,  *$U$  is not observable...*
- But, since its distribution is known, we can “guess” its unobserved value with certain degree of reliability.
- This “guess” is based on a random set  $\mathcal{S} \sim P_{\mathcal{S}}$  in the  $U$ -space.
- There is theory to guide the choice of  $P_{\mathcal{S}}$ .

- Given  $\mathcal{S} \sim P_{\mathcal{S}}$ , push it forward to the  $\theta$  space:

$$\Theta_y(\mathcal{S}) = \bigcup_{u \in \mathcal{S}} \{\vartheta : y = a(\vartheta, u)\}.$$

- The *belief function* is just the containment functional

$$b_y(A) = P_{\mathcal{S}}\{\Theta_y(\mathcal{S}) \subseteq A\}.$$

- Dual is the *plausibility function*

$$p_y(A) := 1 - b_y(A^c) = P_{\mathcal{S}}\{\Theta_y(\mathcal{S}) \cap A \neq \emptyset\}.$$

- Non-additive, i.e.,  $b_y(A) \leq p_y(A)$  for all  $A \subseteq \Theta$ .

## Theorem.

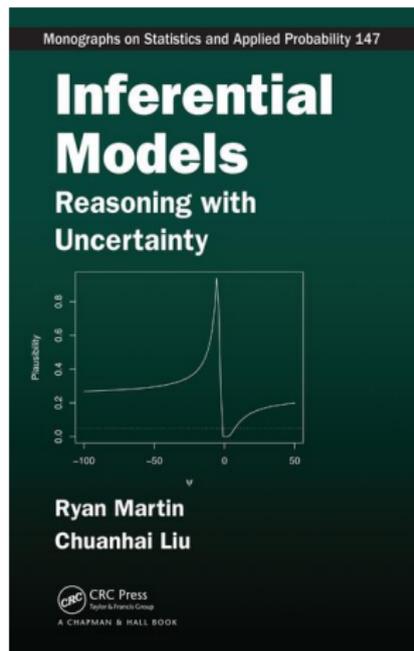
With a “suitable choice of random set  $\mathcal{S} \sim P_{\mathcal{S}}$ ,” the inferential model constructed above satisfies the validity condition.

- Re: “suitable choice of  $\mathcal{S}$ .”
  - Let  $f(u) = P_{\mathcal{S}}(\mathcal{S} \ni u)$ .
  - Need  $f(U) \geq_{\text{st}} \text{Unif}(0, 1)$  when  $U \sim P_U$ .
- This is actually easy to arrange...
- Consequences of validity:
  - “Reject  $H_0 : \theta \in A$  if  $p_Y(A) \leq \alpha$ ” is a size  $\alpha$  test.
  - The  $100(1 - \alpha)\%$  plausibility region

$$\{\vartheta : p_Y(\{\vartheta\}) \geq \alpha\}$$

is a  $100(1 - \alpha)\%$  confidence region.

*Valentine's Day is coming up — there's still time to get the perfect gift for that special someone*



- Let  $Y \sim \text{Bin}(n, \theta)$ , distribution function  $F_\theta$ .
- Construct an inferential model for  $\theta$ :
  - A.  $F_\theta(Y - 1) \leq U < F_\theta(Y)$ ,  $U \sim \text{Unif}(0, 1)$ .
  - P. “Default” random set

$$\mathcal{S} = \{u : |u - 0.5| \leq |\tilde{U} - 0.5|, \tilde{U} \sim \text{Unif}(0, 1)\}.$$

- C. Combine to get<sup>5</sup>

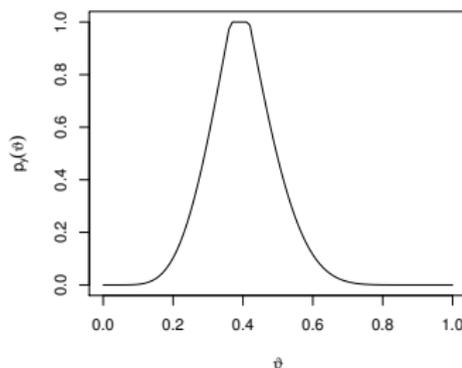
$$\begin{aligned}\Theta_y(\mathcal{S}) &= \bigcup_{u \in \mathcal{S}} \{\theta : F_\theta(y - 1) \leq u < F_\theta(y)\} \\ &= [1 - G_{n-y+1, y}^{-1}(\frac{1}{2} + |\tilde{U} - \frac{1}{2}|), 1 - G_{n-y, y+1}^{-1}(\frac{1}{2} - |\tilde{U} - \frac{1}{2}|)],\end{aligned}$$

- Note this only depends on  $\tilde{U} \sim \text{Unif}(0, 1)$ ...

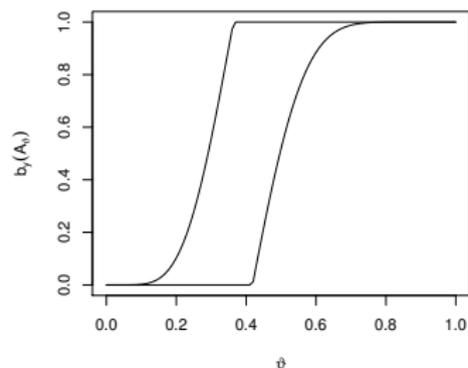
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<sup>5</sup>Use fact that  $F_\theta(y) = G_{n-y, y+1}(1 - \theta)$ , beta distribution.

- Data:  $n = 18$ ,  $y = 7$ .
- Plots of plausibility contour  $p_y(\{\vartheta\})$  and of belief and plausibility of  $A_\vartheta = [0, \vartheta]$ .



(e) Plausibility contour



(f)  $b_y(A_\vartheta)$  and  $p_y(A_\vartheta)$

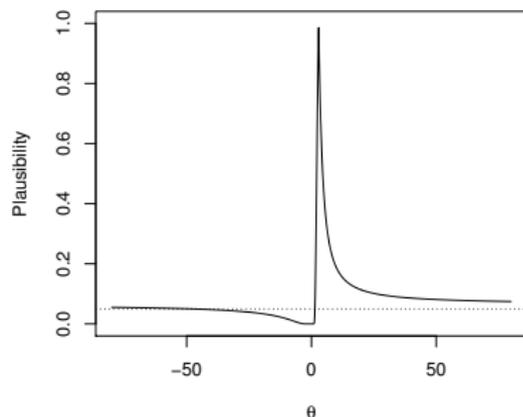
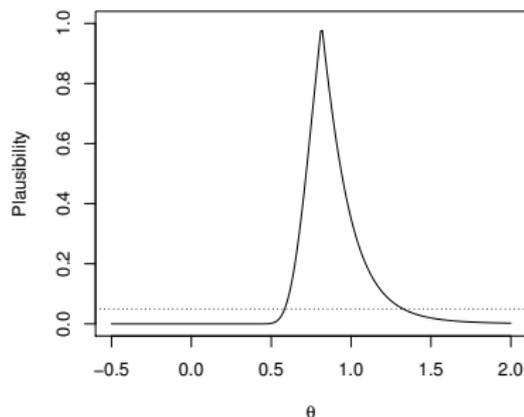
- Let  $Y = (Y_1, \dots, Y_n)$  be iid  $N(\mu, \sigma^2)$ ,  $\theta = (\mu, \sigma)$ .
- Goal is inference on  $\phi = \sigma/\mu$ , coefficient of variation.
- $\mu$  in the denominator makes this a difficult problem.
- Details are too lengthy to present here, so just the highlights:
  - build a (marginal) association:

$$\frac{n^{1/2}\bar{Y}}{\mathcal{S}} = F_{n,1/\phi}^{-1}(U), \quad U \sim \text{Unif}(0, 1).$$

- random set  $\mathcal{S}$  for  $U$  yields a (marginal) inferential model for  $\phi$
- guaranteed valid!

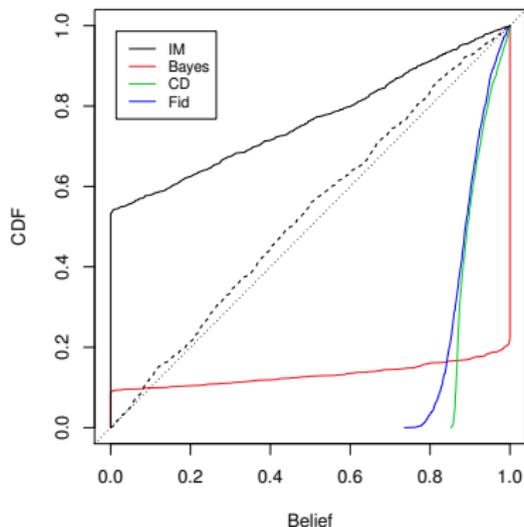
## Normal CV example, cont.

- Simulate data of size  $n = 30$  from  $N(\mu, \sigma^2)$ , with  $\sigma = 1$ .
- Plots of plausibility contour for  $\phi$ ; based on “default”  $\mathcal{S}$ .
- Left:  $\mu = 1$ ; right:  $\mu = 0$ .



## Normal CV example, cont.

- Simulation:  $n = 10$  from  $N(0.1, 1)$ , so that  $\phi = 10$ .
- Consider  $A = (-\infty, 9]$ , which *does not* contain  $\phi = 10$ .
- Compute CDFs of various (marginal) beliefs  $b_y$ .
- Colored lines have false confidence, black lines don't!



- Inferential model construction and validity result is general, but more care is needed to get efficient solutions.
- I hid these steps in the normal CV example above.
- What's often needed is a way to reduce the dimension of  $U$  before introducing the random set to guess its value.
- For example, if we have iid data with

$$Y_i = a(\theta, U_i), \quad i = 1, \dots, n,$$

then the common  $\theta$  and observed  $y_1, \dots, y_n$  imposes a constraint on the unobserved values  $U_1, \dots, U_n$ .

- In other words, *there is a feature of  $U$  that is observed.*

- There exists  $y \mapsto (T(y), H(y))$  and  $u \mapsto (\tau(u), \eta(u))$  such that the above association can be re-expressed as

$$T(Y) = b(\theta, \tau(U)) \quad \text{and} \quad H(Y) = \eta(U).$$

- Key point is that the 2nd expression has no  $\theta$  in it, hence the value of  $\eta(U)$  is observed, even though  $U$  isn't.
  - Don't need to guess  $\eta(U)$ , hence dimension reduction;
  - Observed  $\eta(U)$  helps improve guess of unobserved  $\tau(U)$ .
- How to find  $(T, H)$  and  $(\tau, \eta)$ ?
  - sufficient statistics
  - group invariance
  - *solving a partial differential equation*
  - ...

- Let  $u_{y,\theta}$  be a solution to the equation  $y = a(\theta, u)$ .
- $\eta(U)$  would be observed if  $\eta(u_{y,\theta})$  is constant in  $\theta$ .
- That is, we seek  $\eta$  that satisfies

$$\frac{\partial \eta(u_{y,\theta})}{\partial \theta} = 0.$$

- I don't know about existence in general, but sometimes I can solve it, e.g., method of characteristics.
- Sometimes the solution depends on the parameter, which is bad, but there's a neat *localization* trick to get around this.
- Sorry for the lack of details...

- Location model illustration:  $Y = \theta \mathbf{1}_n + U$ .
- $u_{y,\theta} = y - \theta \mathbf{1}_n$ .
- Goal is to solve

$$0 = \frac{\partial \eta(u_{y,\theta})}{\partial \theta} = \frac{\partial \eta(u)}{\partial u} \Big|_{u=u_{y,\theta}} \cdot \frac{\partial u_{y,\theta}}{\partial \theta}.$$

- Of course, right-most factor is  $\mathbf{1}_n$  in this case.
- Need  $\eta$  such that  $\partial \eta(u) / \partial u$  sends constant vectors to 0.
- Take  $\eta(u) = Mu$  where, e.g.,  $M = I_n - n^{-1} \mathbf{1}_n \mathbf{1}_n^\top$ .
- Then  $U_i - \bar{U} = Y_i - \bar{Y}$  are *observed*.

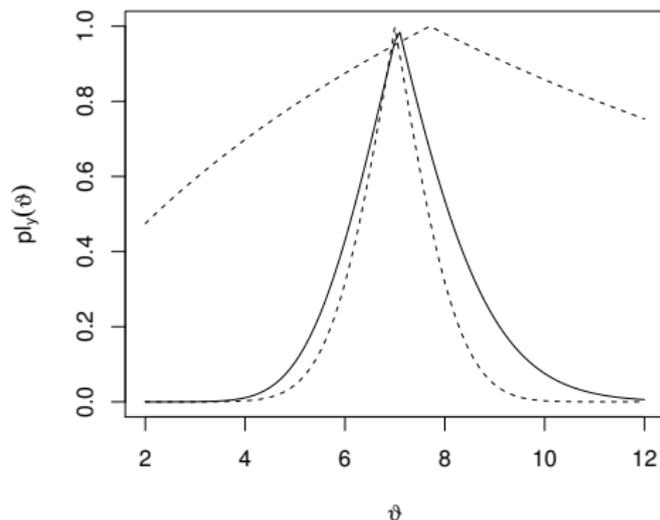
- A variation on the above location model:

$$Y_1 = \theta + U_1, \quad Y_2 = \theta^{1/3} + U_2, \quad U_1, U_2 \stackrel{\text{iid}}{\sim} N(0, 1).$$

- “Non-regular,” no reduction via sufficiency is possible.
- Same PDE as before can be set up, but solving it requires the localization trick I mentioned.
- Skipping details, the (localized) plausibility contour is

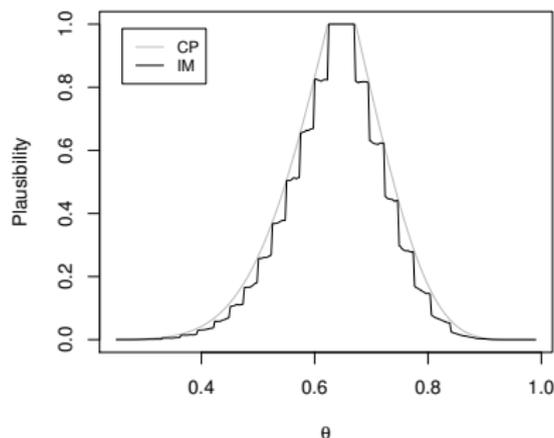
$$p_Y(\{\vartheta\}) = 1 - \left| 2\Phi\left(\frac{y_1 - \vartheta + \frac{1}{3\vartheta^{2/3}}(y_2 - \vartheta^{1/3})}{\left(1 + \frac{1}{9\vartheta^{4/3}}\right)^{1/2}}\right) - 1 \right|.$$

- Cube-root model, true value is  $\theta = 7$ .
- Plot plausibility contours based on individual observations and the one based on solving the PDE.



- Valid inferential models yield nominal confidence regions.
- Interesting reversal:
  - Given a nested family of confidence regions, create a function by stacking up its contours
  - This is a valid plausibility function.
- So, valid belief/plausibility is fundamental to statistics.
  - We always tell students that confidence *isn't* probability, but we never tell them what it *is*.
  - Confidence is plausibility: *a confidence set contains all parameter values that are sufficiently and justifiably plausible based on the observed data.*
  - Similar statements can be made about p-values.

- Complete-class theorem:<sup>6</sup>
  - (basically) for every confidence region there exists a random set such that the corresponding inferential model's is valid and its plausibility region is the same or smaller,
  - and there's an algorithm for constructing the random set.
- Binomial example...



<sup>6</sup>M., arXiv:1707.00486.

- Non-additive beliefs are fundamental to statistics.
- Additive beliefs put user at risk of false confidence and “systematically misleading conclusions.”
- Avoid this risk with suitable non-additive beliefs.
- Main drawback of the proposed approach is that it's not always easy to do — we still don't understand it that well.
- Lots of interesting questions that remain to be answered:
  - marginalization mysteries
  - “optimal” random sets
  - incorporating prior info without losing validity
  - model uncertainty
  - computation!
  - ...

- The peer review system is broken in various ways.
- A particular concern of mine:
  - Feedback and judgment is through the same process
  - “Positive feedback” is good for our careers, but not for science
- Successful reform requires new ideas.
- Harry Crane and I developed a new open-access publication platform, featuring an *author-driven* peer review process.

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