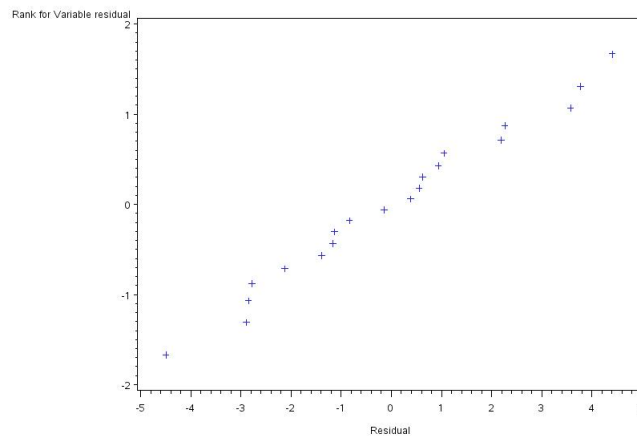


ST 512 Homework assignment #2
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1. Obtain the residuals from a simple linear regression of bodyfat percentage on midarm circumference (x_2) x_3 using the dataset “bodyfat.dat”. Obtain a plot of the empirical quantiles of the residuals (sorted residuals) against the corresponding quantiles from the standard normal distribution. That is, obtain a normal q-q plot of the residuals from the simple linear regression of bodyfat on midarm. In computing the normal quantiles, indicate how you avoided having to compute either the 0th or the 100th quantile of the standard normal distribution.

Normal q-q plot, regression of bodyfat on midarm



PROC RANK with the option NORMAL=VW is one way to obtain the corresponding normal quantiles for any sample of residuals. This option finds the $\{i/(n+1)\}$ normal quantiles for $i = 1, \dots, n$.

```
proc reg data=bodyfat;
    model y=x2;
    output out=predz p=yhat r=residual;
run;
proc rank data=predz normal=vw out=bodyfat2;
    var residual;
    ranks normalscores;
run;
options device=jpeg;
run;
proc gplot data=bodyfat2;
    title "Normal q-q plot, regression of bodyfat on midarm";
    plot normalscores*residual;
run;
```

2. Rao 8.3 and Rao 8.9

(a)

$$\begin{aligned} y_{1+} &= 8.93 \\ y_{2+} &= 13.75 \\ y_{3+} &= 18 \end{aligned}$$

(b)

$$y_{++} = 40.68$$

(c)

$$\begin{aligned} \bar{y}_{1+} &= 2.23 \\ \bar{y}_{2+} &= 3.44 \\ \bar{y}_{3+} &= 4.5 \\ s_1^2 &= 0.91 \\ s_2^2 &= 0.21 \\ s_3^2 &= 0.54 \end{aligned}$$

3. Rao 8.8

(a) $df_{trt} = 3, MS(trt) = 42, SS(E) = 320, SS(TOT) = 446, F_{obs} = 2.625, F(.05, 3, 20) = 3.089$

(b) $t = 4$

(c) $n_i \equiv 6$, since balanced.

(d) No evidence to reject the (null) hypothesis that mean is constant across four treatments.

4. (Rao 12.3a)

The GLM Procedure

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	2	1752.266667	876.133333	6.71	0.0043
Error	27	3527.600000	130.651852		
Corrected Total	29	5279.866667			

5. (a) (Rao 12.5a)

$$Y_i = \beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + \beta_3 x_{i3} + \beta_4 x_{i4} + E_i \text{ for } i = 1, \dots, 20$$

where the indicator variables x_{ij} are defined by

$$x_{ij} = \begin{cases} 1 & \text{seedling } i \text{ receives treatment } j \\ 0 & \text{else} \end{cases}$$

(b) (Rao 12.5b) for $i = 1, 2, 3, 4$ we have $\beta_0 + \beta_i$ as the mean plant height for treatment i and β_0 as the mean for treatment 5.

(c) (Rao 12.5) - estimated regression parameters:

$$\begin{aligned}\hat{\beta}_0 + \hat{\beta}_1 &= 34.02 \\ \hat{\beta}_0 + \hat{\beta}_2 &= 31.90 \\ \hat{\beta}_0 + \hat{\beta}_3 &= 30.84 \\ \hat{\beta}_0 + \hat{\beta}_4 &= 34.31 \\ \hat{\beta}_0 &= 34.29\end{aligned}$$

Successively substituting means leads to

$$\hat{\beta} = \begin{pmatrix} 34.29 \\ -0.27 \\ -2.39 \\ -3.45 \\ 0.02 \end{pmatrix}$$

6. Referring to Rao 12.6b, the contrast described in part (i) is just $\theta_i = \frac{\mu_2 + \mu_3}{2} - \frac{\mu_4 + \mu_5}{2}$ the contrast described in part (ii) is just $\theta_{ii} = (\mu_2 - \mu_4) - (\mu_3 - \mu_5)$ Substitution of treatment means yields the following estimates:

Parameter	Estimate	Standard Error	t Value	Pr > t
i: A-B	-2.93000000	0.52243899	-5.61	<.0001
ii: source x intensity	1.04500000	1.04487798	1.00	0.3331

7. Rao 12.10

(a) Rao 12.10b

$$\mu(x_1, x_2, z) = \beta_0 + \beta_5 z$$

(b) Rao 12.10d - First $H_{0\tau}$ postulates all three adjusted treatment means equal, second $H_{0\tau}$ postulates first two adjusted treatment means equal.

(c) Rao 12.10e

$$\mu(x_1, x_2, z) = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 z$$

(d) Rao 12.10f

$$F = \frac{SS[R]_f - SS[R]_r}{MS[E]_f} = \frac{4899 - 3283}{14.6} = 110.4$$

Wow, that's big. Conclude treatments 1 and 2 different.

8. Rao 12.11d - use the expression below, with $x' = (1, 1, 0, 30)$

$$\hat{\beta}'x \pm t(0.025, 26)\sqrt{x'\hat{\Sigma}x}$$

(where $\hat{\Sigma} = (X'X)^{-1}MS[E]$) to get $52.8 \pm 2.06(1.21)$ or $(50.3, 55.3)$

9. Rao 12.18

- (a) Rao 12.18a

$$Y_i = \beta_0 + \beta_1x_{i1} + \beta_2x_{i2} + \beta_3x_{i3} + \beta_4x_{i4} + \beta_5z_i + E_i$$

where x_{ij} is defined by

$$x_{ij} = \begin{cases} 1 & \text{bag } i \text{ at location } j \\ 0 & \text{else} \end{cases}$$

and $\{z_i\}$ denote initial weights and $\{E_i\}$ denote independent normally distributed error terms with common variance σ^2 for bags $i = 1, \dots, 20$. For $i = 1, 2, 3, 4$, β_i denotes the mean difference between location i and location 5, after controlling for initial weight z . β_5 denotes the mean increase in final weight per unit increase in z for fixed location. β_0 is the intercept of the (population) line describing the association between mean final weight and initial weight in location 5.

- (b) Rao 12.18b - let $\mu_j(z)$ denote mean final weight for initial weight z at location j :

$$\begin{aligned} \hat{\mu}_1(z) &= 2.25 + 1.08z \\ \hat{\mu}_2(z) &= 2.21 + 1.08z \\ \hat{\mu}_3(z) &= 4.15 + 1.08z \\ \hat{\mu}_4(z) &= 3.60 + 1.08z \\ \hat{\mu}_5(z) &= 2.49 + 1.08z \end{aligned}$$

- (c) Rao 12.18c - by evaluating the expressions above at $\bar{z} = 27.7$, we get

$$\begin{aligned} \hat{\mu}_1(\bar{z}) &= 30.15 \\ \hat{\mu}_2(\bar{z}) &= 30.12 \\ \hat{\mu}_3(\bar{z}) &= 32.05 \\ \hat{\mu}_4(\bar{z}) &= 31.50 \\ \hat{\mu}_5(\bar{z}) &= 30.40 \end{aligned}$$

- (d) Rao 12.18d - 90% c.i. for β_1 is $\hat{\beta}_1 \pm t(0.05, 14)SE(\hat{\beta}_1)$ or $-0.24 \pm 1.76(0.5766)$ -0.24 ± 1.02 or $(-1.26, 0.77)$.

- (e) Rao 12.18f - conclusion about $H_0 : \beta_5 = 0$ based on $t_{\beta_5} = 22.75$ which is significantly different from 0 ($p < 0.0001$). Therefore, final weight associated w/ initial weight.
- (f) Rao 12.18g - 95% c.i. given by $\hat{\beta}_5 \pm t(0.025, 14)\sqrt{0.0075MS[E]}$ or $1.08 \pm 2.14(0.0476)$ or $(0.98, 1.19)$.
- (g) Rao 12.18h - 95% p.i. for Y given $x' = (1, 0, 0, 1, 0, 25)$ given by $x'\hat{\beta} \pm t(0.025, 14)\sqrt{MS[E] + x'\hat{\Sigma}x}$ or $31.23 \pm 2.14\sqrt{0.30 + 0.08}$ or 31.23 ± 1.31 or $(29.9, 32.5)$.